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AUTHOR Wasserman, John D.; And Others

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ABSTRACT

IDENTIFIERS

The original Bayley Scales of Infant Development (BSID) have been among the most popular measures of performance and aptitude of infants. In this study, the construct validity of scores on the Behavior Rating Scale of the revised Bayley Scales, the BSID-II, was investigated using national standardization and clinical samples of children ranging in age from roughly 1 to 42 months, and a variety of factor analytic methods. In all, 2,106 children (562 1- to 5-month-old, 503 6- to 12-month-old, and 1,041 13- to 42-month-old subjects) provided data for the analyses. Results indicate that motor performance has an important influence on scores on the BSID-II Behavior Rating Scale, a result that is judged to be consistent with expectations. It is also clear that the structure underlying scale scores becomes increasingly more complex as the samples become more heterogeneous. The least heterogeneous samples are at the youngest age where development is least differentiated. These results offer insight into the integrity of scores from the new Bayley Scales. Seventeen tables present analysis details. (Contains 26 references.) (SLD)



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THE FACTOR STRUCTURE OF THE BEHAVIOR RATING SCALE
OF THE BAYLEY SCALES OF INFANT DEVELOPMENT-II:
CROSS-SAMPLE, CROSS-SECTIONAL, AND CROSS-METHOD
INVESTIGATIONS OF CONSTRUCT VALIDITY

John D. Wasserman

The Psychological Corporation

Bruce Thompson

Texas A&M University and Baylor College of Medicine

Kathleen Matula

The Psychological Corporation

ABSTRACT

The original Bayley Scales of Infant Development (BSID) (Bayley, 1969) have been among the most popular measures of performance and aptitude of infants. In this study we investigated the construct validity of scores on the Behavior Rating Scale of the revised Bayley Scales, the BSID-II. We employed national standardization samples of children ranging in age from roughly 1 to 42 months, and a variety of factor analytic methods.



The original Bayley Scales of Infant Development (BSID) (Bayley, 1969) have been among the most popular measures of performance and aptitude of infants. The initial publication of the BSID prompted a host of investigations regarding the characteristics of intelligence of infants ranging in age from roughly 1 to 42 months. Two parts of the original BSID—the mental scale and the motor scale—have been the primary focus of previous research efforts.

However, a third part of the measure—the Behavior Rating Scale (formerly labelled the Infant Behavior Record or IRB)—has also undergone considerably less clinical and experimental use. As noted by Matheny (1980, p. 1157) with respect to the Behavior Rating Scale, the scale is "considered by Bayley [and others] to provide useful information about infants' developmental status, but it has not received nearly as much attention" as the very thoroughly researched mental and motor scales. The IBR's most widely cited use has been in twins studies (e.g., Freedman, 1965; Goldsmith & Gottesman, 1981; Matheny, 1983), where it has been helpful in shedding important insights into the origins of aptitude and performance. Efforts to understand intelligence and behavior in infants is important to efforts to understand development more generally, and studies of development may ultimately have important implications for the ways we educate youngsters.

The revised Scales, the BSID-II, are currently being released, and the new BSID-II will doubtless spark at least as much research and attendant insight and controversy. The present study was



conducted to explore the construct validity of scores from the new Behavior Rating Scale of the new BSID-II.

Our study was grounded on the philosophical premises that the business of science is formulating *generalizable* insight, and that no one study, taken singly, establishes the basis for such insight. As Neale and Liebert (1986, p. 290) observed:

one study, however shrewdly designed No carefully executed, can provide convincing support for a causal hypothesis or theoretical statement... Too many possible (if not plausible) confounds, generality, and alternative limitations on offered for can be any interpretations observation. Moreover, each of the basic methods of research (experimental, correlational, and case study) and techniques of comparison (within- or between-subjects) has intrinsic limitations. How, then, does social science theory advance through research? The answer is, by collecting a diverse body of evidence about any major theoretical proposition.

In the context of the analytic methods that we employed-factor analyses--Gorsuch (1983, p. 201) made a related observation that, "Factors that will appear under a wide variety of conditions are obviously more desirable than factors that appear only under specialized conditions", e.g., only when certain samples or certain factor extraction or rotation methods are used.

Given our premises, we investigated the structure underlying BSID-II Behavior Rating Scale scores across (a) two different types of samples of subjects, (b) three different age groups of subjects within each sample type, and (c) using both first-order and second-order factor analyses and several factor rotation strategies. Specifically, with regard to sampling, we investigated structure using both BSID-II national standardization samples and samples consisting only of children with identified exceptionalities, and also using the combination of these two sample types. The three age cohorts we considered were: (a) children 1 to 5 months of age, (b) children 6 to 12 months of age, and (c) shildren 13 to 42 months of age. Different though overlapping items are used at each of these three ages.

Our study was conducted to address three research questions. First, what is the first-order factor structure underlying responses to BSID-II Behavior Rating Scale items across the two sample types and the combined sample across the three age groups? Second, what is the second-order factor structure underlying responses to BSID-II Behavior Rating Scale items across the two sample types and the combined sample across the three age groups? Third, do differences in mean factor scores across the two sample types also provide evidence of construct validity of BSID-II Behavior Rating Scale score?

Empirical Research with the Scale on the Original BSID

The structure underlying scores on the original BSID Behavior Rating Scale (formerly labelled the Infant Behavior Record) was



investigated independent by various researchers (e.g., Wolf & Lozoff, 1985; Matheny, 1983; Matheny, Dolan & Wilson, 1974; Sameroff, Seifer & Zax, 1982; and Becker, Lederman & Lederman, 1989). Matheny's work was at the vanguard of these efforts. Based upon analysis of the items that had 5- or 9-point rating scales, Matheny (1983) proposed that three gender- and age-invariant factors underlay BSID Behavior Rating Scale items: (a) Task Orientation, (b) Test Affect-Extraversion, and (c) Activity. This three-factor solution was supported by results in several other studies (Braungart, Plomin, DeFries & Fulker, 1992; Plomin & DeFries, 1985; Kaplan, Jacobson & Jacobson, 1991).

Sameroff, Seifer and Zax (1982), however, extracted five factors in their research with children 4 and 12 months old, and six factors for children 30 months old. Although these factors appeared to overlap with those reported by Matheny, few details of their analyses were reported. Fried and Watkinson (1988) studied infants prenatally exposed to marijuana, cigarettes, and alcohol, and extracted factors similar to those obtained by Matheny. Jacobson and Jacobson analyzed data (1991)Kaplan, continuously-scaled BSID Behavior Rating Scale items completed for low-income African American infants at ages 13 and 25 months. They extracted three factors--Test Affect, Test Attention, and Arousal-that closely resembled Matheny's (1980), notwithstanding some variation in the contribution of individual items to each factor.

Method

Sample



Test publishers are frequently willing to provide researchers with access to standardization sample data when they can be assured that the data will be treated as proprietary information, when thoughtful analytic proposals have accompanied requests for data access. We are grateful to the Psychological Corporation for providing us with access to BSID-II Behavior Rating Scale data for the purposes of the present study.

We were provided with access to national samples of data both for the standardization sample and for a clinical sample of children with diagnosed exceptionalities. Table 1 summarizes the demographic characteristics for the sample of 2,106 children who provided the data for our analyses.

INSERT TABLE 1 ABOUT HERE.

<u>Results</u>

Research Question #1

The study's first research question asked, what is the firstorder factor structure underlying responses to BSID-II Behavior
Rating Scale items across the two sample types and the combined
sample across the three age groups? Many researchers acknowledge
the prominent role that factor analysis can play in efforts to
establish construct validity. For example, Nunnally (1978, p. 111)
noted that, historically, "construct validity has been spoken of as
[both] 'trait validity' and 'factorial validity.'"

Similarly, Gorsuch (1983, p. 350) noted that, "A prime use of factor analysis has been in the development of both the operational



constructs for an area and the operational representatives for the theoretical constructs." In short, "factor analysis is intimately involved with questions of validity.... Factor analysis is at the heart of the measurement of psychological constructs" (Nunnally, 1978, pp. 112-113). We employed both first-order and second-order factor analysis in the present study. We employed principal components analyses for all factor extractions.

Analysts differ quite heatedly over the utility of principal components as compared to common or principal factor analysis. For example, an entire special issue on this controversy was recently published in *Multivariate Behavioral Research*. The difference between the two approaches involves the entries used on the diagonal of the correlation matrix that is analyzed—principal components analysis uses ones on the diagonal while common factor analysis uses estimates of reliability, usually estimated through an iterative process.

The two methods yield increasingly more equivalent results as either (a) the factored variables are more reliable or (b) the number of variables being factored is increased. Snook and Gorsuch (1989, p. 149) explain this second point, noting that "As the number of variables decreases, the ratio of diagonal to off-diagonal elements also decreases, and therefore the value of the communality has an increasing effect on the analysis." For example, with 10 variables the 10 diagonal entries in the correlation matrix represent 10% (10 / 100) of the 100 entries in the matrix, but with 100 variables the diagonal entries represent



only 1% (100 / 10,000) of the 10,000 matrix entries. Gorsuch (1983) suggests that with 30 or more variables the differences between solutions from the two methods are likely to be small and lead to similar interpretations.

with respect to the 562 children aged 1 to 5 months in our study, based on application of Cattell's "scree" test to the eigenvalues (i.e., 5.90, 2.97, 1.40, 1.16, 0.94, etc.) prior to factor rotation (Thompson, 1989), we extracted two first-order factors for these data. We extracted the same number of factors in separate analyses for the standardization (n=387) and the clinical (n=175) samples. Because many items correlated with more than one factor, we rotated the first-order solution obliquely. We employed promax rotation for this purpose. As Gorsuch (1983, p. 191) notes, "because the procedure [promax rotation] gives good simple structure, is easily programmed, and is extremely fast, its popularity is spreading rapidly." Table 2 presents these results.

INSERT TABLE 2 ABOUT HERE.

With respect to the 503 children aged 6 to 12 months, based on application of Cattell's "scree" test to the eigenvalues (i.e., 8.55, 5.21, 1.66, 1.50, 1.31, etc.) prior to factor rotation (Thompson, 1989), we extracted three first-order factors for these data. We extracted the same numbers of factors in separate analyses for the standardization (n=315) and the clinical (n=188) samples. Because many items correlated with more than one factor, we again rotated the first-order solution to the promax criterion.



Table 3 presents these results.

INSERT TABLE 3 ABOUT HERE.

With respect to the 1,041 children aged 13 to 42 months, based on application of Cattell's "scree" test to the eigenvalues (i.e., 8.21, 4.83, 2.17, 1.80, 0.94, etc.) prior to factor rotation (Thompson, 1989), we extracted three first-order factors for these data. We extracted the same number of factors in separate analyses for the standardization (n=639) and the clinical (n=402) samples. Because many items correlated with more than one factor, we again rotated the first-order solution to the promax criterion. Table 4 presents these results.

INSERT TABLE 4 ABOUT HERE.

Research Question #2

We also investigated the structure underlying responses using second-order factor analysis. The study's second research question asked, what is the second-order factor structure underlying responses to BSID-II Behavior Rating Scale items across the two sample types and the combined sample across the three age groups?

Kerlinger (1984) noted that, "while ordinary factor analysis is probably well understood, second-order factor analysis, a vitally important part of the analysis, seems not to be widely known and understood" (p. xivv). Example applications of second-order factor analysis have been reported by Kerlinger (1984), Thompson and Borrello (1986), and by Thompson and Miller



(1981).

Gorsuch (1983) emphasizes that the extraction of correlated factors implies that second-order factors should be extracted. He noted, "Rotating obliquely in factor analysis implies that the factors do overlap and that there are, therefore, broader areas of generality than just a primary factor. Implicit in all oblique rotations are higher-order factors. It is recommended that these be extracted and examined..." (p. 255).

Thompson (1990, p. 575) explains second-order analysis:

Many researchers are familiar with the extraction of
principal components from either a variancecovariance matrix or a correlation matrix. However,
the factors extracted from such matrices can be
rotated obliquely such that the rotated factors
themselves are correlated. This interfactor matrix
can then, in turn, also be subjected to factor
analysis. These 'higher order' factors would be
termed second-order factors.

However, it is important not to try to interpret these secondorder factors without first relating them back to the observed
variables themselves. Interpreting second-order factors only with
reference to the first-order factors has been likened to
interpreting shadows (second-order factors) made by other shadows
(first-order factors) caused by real objects (the actual
variables).

Even some very sophisticated researchers incorrectly attempt



to interpret the second-order factors using the first-order factors. For example, in his review of Kerlinger's (1984) second-order analyses, Thompson (1985) noted that

It is particularly disturbing that the second-order factors are interpreted [by Kerlinger] in terms of the first-order factors. A number of strategies for relating the second-order structure back to the original items have been proposed and would have been appropriate. (p. 430)

As Gorsuch (1983) argued,

Interpretations of the second-order factors would need to be based upon the interpretations of the first-order factors that are, in turn, based upon the interpretations of the variables... To avoid basing interpretations upon interpretations, the relationships of the original variables to each level of the higher-order factors are determined. (p. 245)

Gorsuch (1983, p. 247) suggested that one way to avoid "interpretations of interpretations" is to postmultiply the first-order factor pattern matrix times the orthogonally rotated second-order factor pattern matrix. However, if rotation is used to facilitate interpretation of other structures, it also seems plausible to rotate the product matrix itself to the varimax criterion.

For the purposes of the second-order analyses, we employed



Guttman's (1954) criterion, and extracted all first-order factors with eigenvalues greater than 1.0. These first-order factors were rotated to the promax criterion, the interfactor correlation matrix was analyzed and second-order factors were extracted and rotated to the varimax criterion. First-order factors were then post-multiplied by the second-order factors, as recommended by Gorsuch (1983), and the product matrix was then rotated to the varimax criterion. These analyses were conducted with program SECONDOR (Thompson, 1990). Tables 5 through 7 present these results for the three age groups and the various samples.

INSERT TABLES 5 THROUGH 7 ABOUT HERE.

There is another very intriguing way to interpret second-order factors that also avoids the interpretation of shadows of shadows of real objects. This is the solution proposed by Schmid and Leiman (1957), and explained by Gorsuch (1983, pp. 248-254). This solution "orthogonalizes" the two levels of analyses to each other and also allows interpretation of both levels of analysis in terms of the observed variables. Tables 8 through 10 present the Schmid-Leiman solutions, computed by program SECONDOR (Thompson, 1990), for the data from the children aged 1 to 5 months. It should be noted that the first two columns in Table 8, for example, are also equivalent to the unrotated product matrix that Gorsuch (1983, p. 247) suggested could be interpreted without rotation.

INSERT TABLES 8 THROUGH 10 ABOUT HERE.

Tables 11 through 13 present the Schmid-Leiman solutions for the data from the children aged 6 to 12 months. Tables 14 through 16 present the Schmid-Leiman solutions for the data from the children aged 13 to 42 months.

INSERT TABLES 11 THROUGH 16 ABOUT HERE.

Research Question #3

The study's third research question asked, do differences in mean factor scores across the two sample types also provide evidence of construct validity of BSID-II Behavior Rating Scale scores? To address this question, the varimax-rotated product matrices for the combined samples reported in Tables 2 through 4 were used to create factor scores for each of the three age groups. We then tested the mean differences across the standardization and the clinical samples across each factor for each of the three age groups. These results are reported in Table 17.

INSERT TABLE 17 ABOUT HERE.

Discussion

As Nunnally (1978, p. 298) noted, "one tends to take advantage of chance in any situation [all parametric methods] where something is optimized from the data at hand", as in least squares methods. And as Gorsuch (1983, p. 330) noted, "In factor analysis, one has numerous possibilities for capitalizing on chance." Thus, we were interested in detecting factor structure that was reasonably stable across samples and across analytic methods.



Three precepts guided our interpretation of these results. First, we recognized that item or variable means do not directly affect factor structure. Factors extracted from product-moment correlation coefficients, as in the present study, are "scale-free", i.e., item means do not directly impact results. This is because product-moment correlation coefficients are themselves scale-free. For example, the correlation coefficients between all three pairs of variables (X and Y, X and Z, Y and Z) are all +1.0, even though the means of the variables differ:

 $f X & Y & Z \\ f Jon & 1 & 1 & 3 \\ f Jane & 2 & 2 & 4 \\ Mike & 3 & 3 & 5 \\ f f X & 2 & 2 & 4 \\ f$

This meant that differences in items means across the two samples in a given age group might not necessarily create structure differences across the groups. That is, the structures might differ because relationships among variables differed, but differences in means per se do not yield such differences. If the only differences across samples are developmental delays, then structures will be comparable across groups.

Second, we recognized that restriction of range or variability does attenuate product-moment correlation coefficients, which in turn impacts factor structure (Dolenz, 1992). If subjects in a given sample generally score near the measurement "floor" or "ceiling", then the variability of scores on items will be smaller,

and correlation coefficients among these scores will be attenuated. We expected some of these effects in our samples. For example, clinical subjects may have more homogeneous scores because of developmental delays. More importantly, the combined samples involving both the standardization and the clinical samples were by definition more heterogenous, and therefore their scores were more variable and the structure underlying these scores should theoretically be the most stable and generalizable.

Third, we recognized that factor order within solutions and factor scaling directions were unimportant. With respect to order, a given construct may emerge as Factor I in one sample, Factor II in another, and Factor III in yet a third sample. Small variations in the distribution of factor variance or trace (Thompson, 1989) are not noteworthy; what counts is whether the construct is reasonably stable regardless of ordering across solutions.

With respect to factor scaling, the direction in which a factor is scaled is generally arbitrary. For example, in one data set the variable "handsome" may have a structure coefficient on Factor I of +.9, while "ugly" has a structure coefficient of -.8. In a second sample the signs of the coefficients may be reversed. The construct still remains a measure of attractiveness. We can always legitimately "reflect" any factor by multiplying all the coefficients on the given factor by -1. This is legitimate because in the social science we do not presume any meaningful difference between abstract constructs scaled in different directions. For example, an achievement test can be scores number of right answers



correct, or numbers of wrong answers. Thus, we did not attend to factor order or scaling direction differences in our interpretation.

Research Question #1

The study's first research question asked, what is the firstorder factor structure underlying responses to BSID-II Behavior
Rating Scale items across the two sample types and the combined
sample across the three age groups? With respect to children ages
1 to 5 months, as suggested by the results reported in Table 2, a
Motor Quality factor emerged as Factors I in all three analyses
(standardization only, clinical only, and the combined sample).
Factor II was reasonably similar across the three samples, and we
named the factor, Attention. One noteworthy finding is that Factor
I in the standardization sample was more of a "G" or "General"
factor, as indicated by the structure coefficients reported in
Table 2, and by the higher interfactor correlation coefficient
between the factors for this sample. The factors did not emerge as
identical constructs across samples, but do have identifiable
commonalities.

With respect to the children aged 6 to 12 months, as suggested by the results reported in Table 3, a Motor Quality factor emerged as Factors III, I, and I in the three solutions respectively. A factor we labelled Orientation/Engagement emerged as Factors I, II, and II, respectively. A factor we labelled Emotional Regulation emerged as Factors II, III, and III, respectively. Again, the constructs emerge as related but not identical factors across the



samples.

with respect to the children aged 13 to 42 months, as suggested by the results reported in Table 4, a Motor Quality factor emerged as Factors II, I, and II, respectively. A factor we labelled Orientation/Engagement emerged as Factor III in all three solutions. A factor we labelled Emotional Regulation emerged as Factors I, II, and I, respectively.

Research Question #2

The study's second research question asked, what is the second-order factor structure underlying responses to BSID-II Behavior Rating Scale items across the two sample types and the combined sample across the three age groups? We wanted to also analyze the data with second-order factor analysis, because various levels of analysis give different perspectives on data (Gorsuch, 1983, p. 240). As Thompson (1990, p. 579) explained, "The first-order analysis is a close-up view that focuses on the details of the valleys and the peaks in mountains. The second-order analysis is like looking at the mountains at a greater distance, and yields a potentially different perspective on the mountains as constituents of a range. Both perspectives may be useful in facilitating understanding of data."

Table 5 presented the varimax-rotated product matrix ($F_{18x4} \times F_{4x2} = F_{18x2}$) relating the two second-order factors for the children aged 1 to 5 months through the four first-order factors back to the original 18 variables used at this age level. In the clinical and the combined samples an Attention factor emerged as Factor I while



a Motor Quality factor emerges as Factor II. However, for the standardization sample Factor I was a "G" factor. This result is consistent with the Table 2 first-order finding of a more saturated factor in this sample.

Table 6 presented the varimax-rotated product matrix for the children aged 6 to 12 months. The Motor Quality factor emerged as Factors II, III, and II, respectively. An Orientation/Engagement factor emerged as Factor I in the three solutions. The remaining basically III, respectively) was factor (III) II and uninterpretable in the standardization and clinical samples. However, in the more heterogenous combined sample (which also has the most subjects, i.e., 503), the factor appeared to measure Emotional Regulation.

Table 7 presented the varimax-rotated product matrix for the children aged 13 to 42 months. The Motor Quality factor emerged as Factors II, II, and I, respectively. An Orientation/Engagement factor emerged as Factors III, I, and II, respectively. Factor III for the clinical sample was not readily interpretable. An Emotional Regulation factor emerged as Factors I and III in the standardization and the combined samples, respectively.

The Schmid and Leiman (1957) solutions presented in Tables 8 through 16 provided yet another way to view the data. These solutions present the *unrotated* product matrices (as against the varimax-rotated product matrices presented in Tables 5 through 7) as the first several columns, followed by the first-order factors with all variance present in the second-order product matrices



removed from these first-order matrices. Thus, the residualized first-order factors show what's left of the first-order factors, given the presence of the second-order factors. If the second-order factor perfectly repreduce the variance of a first-order factor, the residualized first-order factor will have a trace of 0.0.

The results presented in these tables are generally consistent with the interpretations presented Parlier. However, Tables 8, 11, and 16 provide some interesting insights, and also illustrate the utility of the Schmid and Leiman (1957) solution. In these three tables the motor factor emerges as Factors I, II, and I, respectively. In all three tables, this factor accounts for the most trace (variance), i.e., 5.42, 5.24, and 5.93, respectively, and represents something of a "G" or "General" factor. In all three solutions, a related first-order factor appears as Factors A, B, and B, respectively.

The second-order Motor factor emerged as a more general activity factor, while the residualized and thus "orthogonalized" first-order factor more narrowly measures movement per se. One implication of these results is that movement saturates the factor space for the test and does so at several levels of analysis simultaneously.

Research Question #3

The study's third research question asked, do differences in mean factor scores across the two sample types also provide evidence of construct validity of BSID-II Behavior Rating Scale



score? Table 17 presented the relevant results. These analyses evaluated differences in mean factor scores across the standardization as against the clinical samples. The varimax-rotated product matrices for the combined samples (presented in Tables 5 through 7) were used to compute these factor scores; these factors involved the most subjects and were most heterogenous, and therefore should yield the most generalizable scoring structures.

As reported in Table 17, statistically significant differences were noted only for Factors II, II, and I, espectively. As indicated in our interpretation of the combined samples results in Tables 5 through 7, these three factors are all the Motor Quality dimension. Thus, the two samples consistently differed on the average on this factor, but did not differ on other dimensions.

Summary

In a practical context, it is important to be able to measure abilities and behaviors of very young children, so that we may be able to identify those who may need and benefit from early intervention. In a scientific context, it is important to develop theory about the nature and the dynamics of aptitude and behavior as regards even very young children. Of course, deriving meaningful measurement of very young children is a daunting task.

Considerable effort has been invested in exploring the constructs measured by the Bayley Scales of Infant Development (BSID) since their development some 25 years ago (Bayley, 1969). The release of revised scales, BSID-II, may facilitate even greater insight regarding dynamics within young children. The present



study focused on one scale from the BSID-II, the Behavior Rating Scale (formerly labelled the Infant Behavior Record on the original BSID). Two general conclusions emerge from our research.

First, it is clear that motor performance is an important influence on scores on the BSID-II Behavior Rating Scale. This dimension has particularly noteworthy influences on the factor structure underlying data on the scale. The factor emerges as a distinct entity across analyses. In some analyses (e.g. the standardization sample for children aged 1 to 5 months) the dimension tends to be a "G" factor that dominates the factor space. In several Schmid and Leiman (1957) solutions, motor dynamics emerge as strong influences at both first-order and second-order levels.

From a construct validity point of view, the question is whether this result is consistent with theoretical expectations. Given the nature of motor behavior and the item pools used on the scale at various age levels, we believe the result is consistent with expectations. Motor behavior is most easily discerned by the observer of very young children, and these behaviors seem conceptually discrete from the other items on the scale. It is also likely that motor quality mediates other aspects of performance on the Behavior Rating Scale. The remaining items on the scale are more abstract in their nature, and therefore the theoretical relationships among these items are less obvious.

Second, it is clear that the structure underlying scale scores becomes increasingly more complex as the samples become more



heterogenous. The least heterogenous samples are at the youngest age, where developmental is least differentiated. And within the sample of children aged 1 to 5 months, the standardization sample is the most homogeneous. Thus, this sample at this age yields a two-dimensional structure in which one factor tends to be a "G" or "General" factor that is most highly correlated with the second factor, as reported in Tables 2, 5 and 8.

One expects more homogeneity in younger and less differentiated samples, and consequently expects less complex factor structure is expected here. One expects more heterogeneity in older more differentiated samples, and this expectation was confirmed. Thus, this result seems favorable with regard to construct validity.

Of course, no one study establishes the construct validity of scores from any measure. It will be important to replicate these results in other samples and across various analytic methods. However, results from relatively large national samples do offer important insights regarding the integrity of scores from the new Bayley Scales.



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Table 1
Sample Demographic Characteristics

Months of Age	Standardi	zation	Clinical		Combined	
or Age	Scandard	LZACION	CIIIIICAI		COMBINED	
1-5	n=387 Female Male		175 Female Male		562 Female Male	49.1% 50.9%
	Major. Nonmaj.		Major. Nonmaj.			73.0% 27.0%
6-12	n=315 Female Male		188 Female Male		503 Female Male	51.6% 48.4%
			Major. Nonmaj.			74.3% 25.7%
13-42	n=639 Female Male	49.5% 50.5%	402 Female Male	46.3% 53.7%	1041 Female Male	48.5% 51.5%
	Major. Nonmaj.	70.6% 29.4%	- .	70.8% 29.2%		70.6% 29.4%

Note. "Major." is majority race, while "Nonmaj." is other.



Table 2
Promax-Rotated First-Order Structures for Ages 1 to 5 Months

Combined (n=562) PATTERN MATRIX: 1 .07056 .6285 2 .669091398 3 .06770 .7273 4000532426 500729 .3669	6 .78865 .0 7 .00405 .7 800539 .5 9 .06859 .8 10 .04760 .6 11 .02277 .7 1232285 .3 1319004 .3 14 .85592 .0 15 .90571 .0 16 .89476 .0 17 .765111	FACTOR r MATRIX: 1 1.0000033138 233138 1.00000 STRUCTURE MATRIX: 113774 .60520 2 .7154236152 317334 .70496 4 .0798724243 512891 .36941 6 .7827924243 723619 .72365 819181 .56434 919181 .56434 919181 .56434 1122244 .73243 1242391 .41195 1332207 .46141 14 .8534027604 15 .8629119767 17 .8030936815
n=175) ATRIX: 106 .5333 1110244 615 .7672 9372027	5290 .01270 4531 .71463 4478 .64837 1202 .82980 0197 .71054 1782 .72270 5559 .43348 5143 .47559 9167 .01951 1998 .02458 2157 .05397 382602299	MATRIX: 0000020528 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.0000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.00000 0528 1.0000000 0528 1.000000000000000000000000000000000000
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n (n=3 [X: 8999 8719 3898	10000000000000000000000000000000000000	ATRIX: 0042358 58 1.00000 MATRIX: 96 .80669 4878390 37 .58420 9031398 77 .37670 11830165 60 .78553 74 .37751 116 .70289 30 .53521 110 .53521 110 .53521 110 .53521 110 .27884 111 .27884 111 .27884 11225896 113 .25896 114 .27884
dardizat 1ERN MAT 12002 2078 2753		TOR r M 1.0000 1 1.00
Stan PAT 1 2 3 3	test 8 118 119 12 111 112 114 114 114 115 116 11	F B
Variable Predominant state Lability of arousal Positive affect Negative affect	Scothability when upset Hypersensitivity to test Energy Adaptation to change of tes Interest in test materials Exploration of objects Orientation to examiner Gross-motor movement Control of movement Hypotonicity Hypertonicity Tremulousness Slow and delayed movement Frenetic movement	

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Table 3
Promax-Rotated First-Order Structures for Ages 6 to 12 Months

bined (n	ALIEKN MAIKIA: 1 - 10190 -51049 - 2364	. 1013 . 01101 . 001300	SACO: SALCO: KIDOO:	.0632446441639.	2079:- C107: T67IO:-	14968 .13/i3 .2002	20c1.= /2110. cls18.	.00111 .800550233	.U611913646	2/50: 5686/: 01150:	0 USB3.3 . 1862.1 USB3	1001.	2 - 100000 - 100000 - 100000 - 1000000 - 100000000	03044 71475 20132	2522	6 67618 13553 - 1791	7 - 04548 11008 5696	13762 .43338	9 7779	0 - 05239 .22389 .4107	1 - 03393 18955 5504	2 03907 . 22817 . 4195	3 87049 - 05276 .0439	4 .85752 .01737 .0971	5 94856 02950 0921	6 .76974 31360 .1794	7 .90558 .079860350	8 .84758 .167671002	TOR r MATRIX:	1.0000018108 -: - 18106 1.00000 -:	23709 .45421 1.0000		UCTURE MATRIX: 13827	88782 19704 2011	05685 .51015 .3532	. 6	22340 .25789 .3039	.848/3 20405 3363	
(n=188)	PATTERN MATRIX:	106291 .484562322	2 .89355 .044761417	3 .03506 .48019 .1628	400786 .239486490	529365 .11609 .1794	6 .87791 .017931213	7 .02367 .79075 .0477	8 .0497119282 .7125	9 .04951 .74889 .1031	1009663 .7 859 1 .0550	11 .72076 .16214 .1403	12 .07421 .31280 .6007	1309085 588450 3546	14 .05550 .6377	15 .82583033460303	19389: 21020:- 82847: 61	17 - 09689 . I /028 06	PCO7: /5//#: /0##T: BT	BOCK CECC OFFICE OF	ASSA TILLI OCCO	21 - 02300 . 12121	1440 - 57500 Acces 52	77/0 18/00: #06/0: 57	75. CD170 COEVO 20	1000 TOTAL 20186 07	25 125 1202: 02128: 02 27 181 06775 0143	28 .88424 .04327 .022	TOR r MATRIX:	1 1.00000141761538	3 - 1538C .43795 1.00000		STRUCTURE MATRIX:		2 - 05806 - 54655 3677	4 .05802043665	533771 .23632 .2754	6 .89402159652484	
Standardization (n=315)		.027341853	54048 .00350 .1836	.04892 .1059	03470 .65695 .2012	.34367 -, 120300985	11198 .535000959	.80831 .07486 .0030	8 .16962567250439	9 .6916306925 .0822	.73156 .04632 .1706	.07116 .833581524	.4398134146 .0986	.60129 20774 .0850	.7600610793 .0724	38169 .24366 .0240	16 .07580 .75364 .1377	46361 .0753	.45814 04017 .0816	. 24690 65808 .0043	.09471 .00462 .7203	.1403012530 .6290	.0933202754 .6764	10538154706644	01819 .005716402	5 .07799 .03928	6441321951/5223	27 .1/248 .4656546561 28 .36527 .6827623888	FACTOR r MATRIX:	.0000033699 .2881	2 33699 1.00000 25315	0000:T GIEGZ:- 91887:	TRIX:	.52037122200242	48872 .13913 .02/0	3 .1981 - 14554 .23077 1 - 19811 - 61771 .02489	35580 - 21116 .0309	1991 .597022636	
	Variable	imo	a roll	0 t t t t t t t t t t t t t t t t t t t		Negacive allect			Elleryy Adamtation to change of test.	rials	ith tasks	Exploration of objects	Attention to tasks		toward tasks		Frustration with inability		- 17	Cooperation	Gross-motor movement	Fine-motor movement	Control of movement	icit	Hypertonicity	Tremulousness	Slow and delayed movement	Frenetic movement Hyperactivity						Predominant state			Negative affect	SCOCHABILITY WHEN WESEL HVDersensitivity to test	

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Orientation to examiner 17	٠	.43529569	41	699	17	986	020	941	7.7	400	-	300	
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FIGURE TO MOVEMENT.	٠.	. י	14 -	3064	28		07212	09430	28	84098	03131 -	.2250	
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Table 4 Promax-Rotated First-Order Structures for Ages 13 to 42 Months

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Standardization	.10843	70532	.53648	19679.	76460	38006	.30153	-,12375	.71650	.62991	.53632	04915	56431	.55883	.07335	.76473	.03792	.11569	.06792		.04745	536	12	636	71785	OR F MATRIX	ŏ	22303	.29297	מת ממוזשטוומשט		5003.	54706	67712	90960	.75839	.55570	.49232	
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	variable Positive affect	Negative affect	ďη ι	Hypersensitivity to test		٠. ١	7	intriactive with the base		Accellation to tasks boxesistence in attempting	1	5 4 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	Frustration with inability			Coneration	GOOFGE GOILON MOVEMENT	Fine-motor movement		cit	Hypertonicity	Tremulousness	Slow and delayed movement		Hyperactivity									Sootnablity when upset	בוופדרדאדוא כם	Energy Masstation to change of test.	als		

(3)
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.65622 .32469 .44429 .42520 .21872 .21872 .22905 .22905 .69624 .71558 .69410 .25176 25176
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36483 2134824 2134824 2134839 2134339 214530 214530 210558 205545 205545 205545 205545 205545
222 222 223 223 223 223 223 223 233 243 253
.46598 .27751 .23190 .23190 .30281 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13848 .13868 .10552 .10552
.50324 .78675 .76256 .86173 .14595 .14595 .93249 .09124 .19558 .19558 .19558 .19558 .19559 .17207
17377 00260 08260 08260 05239 29119 13555 18708 26589 26589 92462 92462 91658
110 111 113 114 114 115 116 117 117 117 117 117 117 117 117 117
25843 21066 19037 16363 108247 18488 13589 62611 71309 73720 73720 73720
.12257 .71256 .71978 .67968 52459 55766 .81237 .20021 .28646 23266 04845 15181 1525
0 1 1 1 1 1 1 1 0 0 1 1 1 1 0 0 1 1 1 1
Exploration of objects Attention to tasks Persistence in attempting Enthusiasm toward tasks Fearfulness Frustration with inability Orientation to examiner Social engagement Cooperation Gross-motor movement Fine-motor movement Fine-motor movement Tremulousness Slow and delayed movement Frenetic movement

32

Table 5 Varimax-Rotated Product (F_{VxF} x F_{FxS} = F_{VxS}) Matrices and h^2 for Ages 1 to 5 Months

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11 	
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Variable 1 Predom State 2 Labil Arousa 3 Posit Affect 4 Negat Affect 5 Soothability 6 Hypersensiti 7 Energy 8 Adapt Change 9 Interest Tes 10 Explor Objec 11 Orient Exami 12 Gross-motor 13 Control Move 14 Hypotonicity 15 Hypertonicit 16 Tremulousnes 17 Slow Delayed 18 Frenetic Mov	
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11 - 205 - 223 - 1444 - 1462 - 221 - 252 - 252 - 255 - 163 - 163 - 163 - 163 - 163 - 163 - 163 - 163 - 179 -	
88. 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	
Standardization (n=387) Variable 1 Predom State .51 2 Labil Arousa51 3 Posit Affect .17 5 Soothability .16 6 Hypersensiti23 7 Energy .6 Hypersensiti23 8 Adapt Change .35 9 Interest Tes .64 10 Explor Objec .65 11 Orient Exami .52 12 Gross-motor .54 13 Control Move .54 14 Hypertonicity61 15 Hypertonicit57 16 Tremulousnes43 17 Slow Delayed67 18 Frenetic Mov43	

Table 6' Varimax-Rotated Product (F_{VxF} x F_{FxS} = F_{VxS}) Matrices and h^2 for Ages 6 to 12 Months

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111	Ý O
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Combined (n=503) Variable 2 Labil arousa 3 Posit Affect 4 Negat Affect 5 Soothability 6 Hypersensiti 7 Energy 8 Adapt Change 9 Interest Tes 10 Initiative 11 Exploration 12 Atten Tasks 13 Persistence 14 Enthusiasm 15 Fearfulness 16 Frus Inabili 17 Orient Exami 18 Soc Engageme 19 Cooperation 20 Gross-motor 21 Fine-motor 22 Control Move 23 Hypertonicity 24 Hypertonicity 24 Hypertonicity 25 Tremulousnes 26 Slow Delayed 27 Frenetic Move 27 Frenetic Move 28 Slow Delayed 27 Frenetic Move 27 Frenetic Move 28 Slow Delayed 27 Frenetic Move 28 Slow Delayed 27 Frenetic Move 28 Frenetic Move 29 Frenetic Move 20 Frenetic Move 20 Frenetic Move 20 Frenetic Move 20 Frenetic Move 21 Frenetic Move 21 Frenetic Move 22 Frenetic Move 21 Frenetic Move 22 Frenetic Move 21 Frenetic Move 21 Frenetic Move 22 Frenetic Move 21 Frenetic Move 22 Frenetic Move 23 Frenetic Move 24 Frenetic Move 25 Frenetic Move 26 Frenetic Move 27 Frenetic Move 27 Frenetic Move 27 Frenetic Move 28 Frenetic Mov	28 Hyperactivit Trace
	5.78
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Variable 1 Predom State 2 Labil arousa 3 Posit Affect 4 Negat Affect 5 Soothability 6 Hypersensiti 7 Energy 8 Adapt Change 10 Initiative 11 Exploration 12 Atten Tasks 13 Persistence 14 Enthusiasm 15 Fearfulness 16 Frus Inabili 17 Orient Exami 18 Googeration 20 Gross-motor 21 Fine-motor 22 Control Move 23 Hypotonicity 24 Hypertonicity 25 Slow Delayed 27 Frenetic Mov	8 Hyperactiv race

Table 7 Varimax-Rotated Product (F_{VxF} x F_{FxS} = F_{VxS}) Matrices and h^2 for Ages 13 to 42 Months

. 426 485	. 291	.611	.501	0/5.		250.	1,00	, 50.	070.	740			.618	.442	737	.429	2/4.	000.	. 20.	. 49 B		4/6	4.0	•	14.39
I h² 315 .686	476	.290	2		47.C.	/1#		57.1	703.	701	.067	.328	720	243	826	.062	035	.011	011	.019	040	020	106	002.	4.8b I
II II .571	.093	017	.703	800	4/5	. 63.5	79/	117.	378	499	244	.059	.312	.618	. 228	.399	.405	384	037	020	031	9	.035	٦.	3./4
I 031	.236	726	.077	.072	.104	2.35	.054	971.	.172	.011	669	619	035	.033	.049	.516	.553	.598	769		9		798	. 7	5.79
(n=562) Affect	hity	ensiti		hange	it Tea	ive	tton	Tasks	ence	Lasm	lness	abili	Exami	Engageme	ation	notor		Move	nicity	uicit	onanea		C Mov	Stivit	
+ 4	a a	Hypersensit	Energy	Adapt Change	Interest	Initiative		Atten 1	Persistence	Enthusiasm	Fearfulnes	Frue Inabili	Orient Exami	Soc End	Cooperation	Gross-motor	Fine-motor	Control	Hypotonicity	Hypertonicit	Tremulousnes	Slow D	Frenetic	Hyperactivit	ce Ce
Combined Variable 1 Posi			n n		7			_			13		15		-		19	_			23		ហ		Tra
h ² .507	.200	.646	.297	.345	.620	. 689	.404	. 595	.637	.743	.727	.436	.624	. 508	.650	.755	. 789	, 754	.822	.767	.821	ω	9	. 523	5.65
113	.252	191	.296	.015	.098	.483	.184	.200	.345	.057	268	267	130	113	.070	.864	.875	.837	029	.078	142	.113	072	159	3.10 1
I III 273	152	.780		.179	160	വ	223	.069	015	.010	908.	.602	.051	333	.135	080	101	193	.903	.867	.895	~			6.58
648 1	.338	031	0	. 559	. 765	. 658	. 566	.742	.720	.860	073	045	TTT.	.620	.792	.032	.120	.126	070	097	014	036	.003	058	5.97
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Table 8
Schmid and Leiman (1957) Solution for
Standardization (n=387) Sample for Ages 1 to 5 Months

Variable	I	II	A	В	С	D	h ²
1 Predom State	.576	063	.057	.660	.056	005	.778
2 Labil Arousa		.086			.000	.010	.770
3 Posit Affect	.648	.018	.023	.082	.455	.023	.634
4 Negat Affect	027	.819	070	096	.021	.101	.696
5 Soothability	.276	406	.134	068	.350	034	.387
6 Hypersensiti	365	.481	.185	.089	138	.053	.429
7 Energy	.692	049	064	.533	.116	001	.783
8 Adapt Change	.459	359	098	124	.317	030	.466
9 Interest Tes	.686	083	016	.271	.322	.004	.654
10 Explor Objec	.650	.099	100	.157	.303	.026	.559
11 Orient Exami	.631	342	049	.145	.318	028	.640
12 Gross-motor	.497	.009	366	106	.113	.006	.405
13 Control Move	.568	021	422	.021	.031	001	.502
14 Hypotonicity	607	098	.367	.096	192	021	.560
15 Hypertonicit	569	091	.508	.090	032	013	.601
16 Tremulousnes	464	.065	.399	219	.155	.015	.450
17 Slow Delayed	679	061	.471	078	038	009	.695
18 Frenetic Mov	479	.101	.516	004	.113	.017	.519
Trace	5.42	1.38	1.44	1.42	.85	.02	10.53

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 2 columns represent the second order factors. The next 4 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).

Table 9
Schmid and Leiman (1957) Solution for
Clinical (n=175) Sample for Ages 1 to 5 Months

Va	riable	I					D	h ²
1	Predom State	.569	.004	011	.519	215	.020	.640
2	Labil Arousa	064	.825	.202	040	072	077	.739
3	Posit Affect	.693	033	.022	.368	.233	.017	.671
4	Negat Affect	426	042	.017	.074	.158	.575	.544
	Soothability	.314	068	030	.242	151	064	.189
	Hypersensiti	034	.854	.208	.013	093	051	.785
	Energy	.631	186	016	.393	.221	.102	.647
	Adapt Change	.706	022	.016	.024	.276	407	.741
	Interest Tes	.781	141	.004	.237	.370	132	.840
	Explor Objec	.580	097	.027	.175	.472	.030	.601
	Orient Exami	.800	025	.004	. 195	.134	364	.829
	Gross-motor	.327	441	054	113	.599	016	.676
_	Control Move	.317	271		084	.645	.053	.600
	Hypotonicity	080		.220	045	.045	.006	.773
15		085		.226	012	.028	.041	.818
	Tremulousnes	006		.225		068	048	.885
17		139		.208		.080	.055	.685
	Frenetic Mov	065	.873	.223		030	.087	.830
	ace	3.81	5.42			1.45		12.49

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 2 columns represent the second order factors. The next 4 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).

Table 10
Schmid and Leiman (1957) Solution for
Combined (n=562) Sample for Ages 1 to 5 Months

Variable		I	II	A	В	C	D	h ²
1 Predo	m State	.607	005	002	.585	113	039	.724
2 Labil	Arousa	231	.637	.178	342	.134	057	.630
3 Posit	Affect	.664	123	.021	.209	.319	097	.612
4 Negat	Affect	414	082	.017	.006	.227	.497	.477
	ability	.466	021	007	033	.012	382	.365
	sensiti	121	.746	.195	.017	011	.062	.613
	У	.650	156	006	.510	.132	.047	.727
	Change	.587	129	.005	038	.272	288	.519
_	est Tes	.737	130	.018	.315	.288	082	.748
10 Explo		.555	186	.024	.194	.440	.066	.578
11 Orien	-	.761	110	.003	.173	.190	300	.752
12 Gross		.256	483	059	099	.460	.072	.529
13 Contr	ol Move	.317	385	025	084	.513	.070	.525
14 Hypot		120	.830	.210	.010	059	.016	.751
	tonicit	079	.868	.221	.056	061	.021	.816
16 Tremu		029	.840	.223	054	.024	052	.762
17 Slow	Delayed	215	.778	.186	021	121	.021	.702
18 Frene	-	046	.859	.225	.071	019	.040	.798
Trace		3.70	4.96	.30	.96	1.09	.61	11.63

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 2 columns represent the second order factors. The next 4 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 11 Schmid and Leiman (1957) Solution for Standardization (n=315) Sample for Ages 6 to 12 Months

Variable	I	II	III	A	В	С	D	E	F	h ²
1 Predom State	.128	.133	882	.056	.012	019	197	011	.016	.855
2 Labil arousa	057	169	.867	068	018	.046	.195	009	058	.832
3 Posit Affect	.665	.208	.077	023	060	.042	002	457	024	.706
4 Negat Affect	169	226	.155	008	095	.410	.020	085	.223	.339
5 Soothability	.483	.085	056	068	.043	068	027	303	088	.355
6 Hypersensiti	097	562	.132	030	.072	.382	.038	.079	051	.505
7 Energy	.547	.249	426	.374	.000	.082	052	.006	130	.709
8 Adapt Change	.233	.467	.057	.131		371		074	012	.440
9 Interest Tes	. 409	.470	257	.404	009	037			.032	.621
10 Initiative	.468	. 425	216	.482	056	.060	.000		042	.690
11 Exploration	.523	.289	187	.381	050	.137		053		.564
12 Atten Tasks	.218	.608	063	.433	.015	221	.018	.075	.069	.668
13 Persistence	.368	.531	137	.453		120	.007	.036	.016	.658
14 Enthusiasm	.537		262		007	_			064	.760
15 Fearfulness	649	045	.161	050	.035		003	.020	.443	.658
16 Frus Inabili		356			066	.496		002	.118	.406
17 Orient Exami	.584		.024			242			374	.629
18 Soc Engageme	.622	.275	.141			024		474	.031	.709
19 Cooperation	.377	.585	. 104			426			035	.723
20 Gross-motor	.169	.580	009			016		075		.608
21 Fine-motor	.179					095		040		.664
22 Control Move	.033	.653			427	050				.641
23 Hypotonicity	251					117		024		.449
24 Hypertonicit		385			. 485			012		. 444
25 Tramulousnes	_	448								.386
26 Slow Delayed		412				137				.517
27 Frenetic Mov		637								.650
28 Hyperactivit	. —	548							014	.662
Trace	3.85	5.24	2.15	1.43	1.50	1.45	.09	.60	.53	16.85

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 6 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 12 Schmid and Leiman (1957) Solution for Clinical (n=188) Sample for Ages 6 to 12 Months

Variable	I	II	III	A	В	С	D	E	F	h ²
1 Predom State	.119	.761	092	004	.157	.000	039	066	.349	.754
2 Labil arousa	182	.073	.897	.052	.000	026	085	.032	.029	.856
3 Posit Affect	.559	.201	.026	.001	.107	022	012	.557	053	.678
4 Negat Affect	382	.365	.044	.000	019	492	.066	.102	.034	.539
5 Soothability	.292	327	307	017	.160	.054	054	.114	168	.359
6 Hypersensiti	174	.002	.886	.051	013	033	089	.074	013	.832
7 Energy	.624	.254	035	.001	.472	009	017	.124	.059	.697
8 Adapt Change	.501	110	.010	.002	126	.500	047	.193	.016	.568
9 Interest Tes	.634	.108	016	.003	.513	.015	005	.044	.014	.680
10 Initiative	.631	.072	162	006	.561	047	.031	008	007	.747
11 Exploration	.479	.271	.030	.004	.478	152	028	.148	.031	.579
12 Atten Tasks	.742	127	013	.004	.327	.341	.080	095	.014	.806
13 Persistence	.748	017	170	005	.458	.116	.119	052	003	.829
14 Enthusiasm	.726	.066	024	.003	.471	.161		001	.038	.782
15 Fearfulness	194	~.236	.824	.048	.078	038	056	105	085	.803
16 Frus Inabili	017	226		.046	.070		017	049	079	.678
17 Orient Exami	.631	.225	123	006	072	.218	.095	.416	.042	.700
18 Soc Engageme	.580	.194	.126	.008	.129	.015	.006	.492	035	.650
19 Cooperation	.740	149	088	001	.122	.456	.033	.085	002	.809
20 Gross-motor	.651	.217	057	.000	.007	065	.588	.055	.033	.827
21 Fine-motor	.729	073	085	001	.084	.090	.518	052		.831
22 Control Move	.682	.045	091	002	.064	032	.544		032	.779
23 Hypotonicity	109	006	.883	.051	030	075			027	.806
24 Hy ertonicit	.024	.143	.832	.050	029	021	.148	009		.743
25 T _mulousnes	028	.055	.937	.055	003	.053	034		.029	.893
26 Slow Delayed	066	136	.822	.048	133				053	.739
27 Frenetic Mov	024	.093	.922	.054	.002	003	.040	.025	.036	.865
28 Hyperactivit		015								.782
Trace	6.69	1.37	7.72	.03	1.73	•97	1.02	.89	.19	20.61

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 6 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).

Table 13
Schmid and Leiman (1957) Solution for
Combined (n=503) Sample for Ages 6 to 12 Months

Variable	I	II	III	A	В	С	D	E	h^2
1 Predom State	.094	117	.179	011		068	091	043	.218
2 Labil arousa	130	.869		.103			002	.024	.788
3 Posit Affect	.096	003	.807	.004		041	.006	.272	.739
4 Negat Affect	419	.126	.068	002	030	505	.116	.054	.469
5 Soothability	.086	221	.415	019	.013	.117	091	.140	.270
6 Hypersensiti	248	.825	036	.095	019	087	028	.021	.761
7 Energy	.345	073	.535	.001	.512	.007	059	.037	.678
8 Adapt Change	.525	075	.116	.007	030	.460	.013	.020	.507
9 Interest Tes	.451	042	.472	.004	.475	.038	.019	.017	.656
10 Initiative	.467	115	.472	006	.499	0 07	.047	.008	.706
11 Exploration	.271	.018	.560	.007	.451	105	.011	.066	.607
12 Atten Tasks	.717	054	.177	.008	.323	.305	.098	065	.759
13 Persistence	.634	142	.338	005	.411	.170	.080	027	.740
14 Enthusiasm	.540	070	.486	.004	.479	.129	.005	.016	.779
15 Fearfulness	144	.778	205	.090	020	038	.019	047	.681
16 Frus Inabili	132	.705	019	.079	.050	151	.060	002	.551
17 Orient Exami	.445	151	.410	007	001	.261	.066	.120	.475
18 Soc Engageme	.138	.064	.791	.012	.050	014	.022		.723
19 Cooperation	.656	128	.286	.002	.064	.441	.056		.732
20 Gross-motor	.626	035				056		.026	.661
21 Fine-motor		047						005	.792
22 Control Move		026				027		007	.678
23 Hypotonicity	110			.101	029			.008	
24 Hypertonicit	.014	.833			.015			007	.711
25 Tremulousnes	045						028		
26 Slow Delayed	087	.748	227					017	
27 Frenetic Mov	110	.889	.002	.105					
28 Hyperactivit	127					026			
Trace	4.50	6.98	3.37	.10	1.66	.97	.71	.20	18.48

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 5 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 14
Schmid and Leiman (1957) Solution for
Standardization (n=939) Sample for Ages 13 to 42 Months

Variable	I	II	III	A	В	C	D	E	h ²
1 Posit Affect	.306	023	.611	004	.034	.111	.156	.014	.506
2 Negat Affect	669					.001	068	.092	.583
3 Soothability	.565	.032	.105	.327	.025	.042	014	042	.443
4 Hypersensiti	666		028	234	011	.001	099	124	.581
5 Energy		170	.707		027	.158	033	.030	.569
6 Adapt Change		~.059	019	.228	.017	022	.206	.009	.667
7 Interest Tes		144	.523	006	031	.067	.325	006	.698
8 Initiative		196	.533			.064	.337	.005	.667
9 Exploration		197		077		.130	.120	.018	.591
10 Atten Tasks		109	.146			020	.383	.027	.755
11 Persistence	.702		.275		013	.006		016	.7 3 3
12 Enthusiasm		053	.485	.058	.001	.060	.346	033	.823
13 Fearfulness	245	965	576	161		141	.062	.005	.445
14 Frus Inabili	564	.154	.048			014		023	.478
15 Orient Exami	.672	.013	.399	.221	.011	.075	.170	066	.699
16 Soc Engageme		087	.607	004	.084	.115	.100	.131	.509
17 Cooperation		034		.194	.016			023	.810
18 Gross-motor	.181			037	.060	.018	026		.561
19 Fine-motor	.251	611	.129	059	.016				.638
20 Control Move		672					058		
21 Hypotonicity	.066			.032	.361		016		
22 Hypertonicit	002	.790	062	.013			049		
23 Tremulousnes	.020	.565	032	045		008			
24 Slow Delayed	073	.622	334	.015	.084	068		282	
25 Frenetic Mov	475	.514	.155	173	.056			240	
26 Hyperactivit	610	.275	.396	124	.025	.115	215	130	.697
Trace	6.24	3.66	3.52	.64	.39	.13	.93	.70	16.20

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 5 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 15
Schmid and Leiman (1957) Solution for
Clinical (n=402) Sample for Ages 13 to 42 Months

٧a	riable	r	II	III	A	В	C	α	h ²
1	Posit Affect	.647	277	107	036	.116	011	.355	.648
	Negat Affect	378	310	029	029	361	.001	.189	.408
3	Soothability	.337	148	.255	052	.178	.011	008	.235
4	Hypersensiti	029	.777	204	.236	027	006	.064	.706
5	Energy	.401	215	.301	.018	072	.016	.382	.449
6		.560	.178	.012	.003	.393	004	093	.508
7	Interest Tes	.764	161	.102	013	.217	.000	.293	.753
8	Initiative	.657	143	.487	.002	.165	.024	.242	.775
9	Exploration	.565	221	.189	.020	021	.008	.452	.609
10	Atten Tasks	.742	.070	.200	.009	.360	.006	.077	.731
11	Persistence	.719	011	.346	.005	.298	.015	.127	.742
12	Enthusiasm	.860	.008	.058	.016	.322	002	.235	.902
13	Fearfulness	070	.802	282	.215	.038	011	050	.777
	Frus Inabili	043	.598	277	.195	080	011	.126	.496
15	Orient Exami	.778	.046	129	.001	.359	013	.137	.772
16	Soc Engageme	.619	337	106	035	.046	010	.427	.694
17	Cooperation	.792	.134	.069	.016	.412	002	.060	.823
18	Gross-motor	.030	064	.866	.019	054	.050	.010	.761
19	Fine-motor	.118	086	.876	.003	.011	.049	012	.792
20	Control Move	.124	178	.841			.047	.022	.758
21	Hypotonicity	067	.903	045	.256	.022	.004	045	.890
22	Hypertonicit	094	.869	.062	.248			060	.833
23	Tremulousnes	011		158	.249	.052	003	030	.887
24	Slow Delayed	033	.878	.097	.248	.047	.011	068	.849
25	Frenetic Mov	.006		087	.256	.022			.835
26	Hyperactivit	055	.701	171	.215	043	004	.059	.575
Tr	ace	5.96	6.54	3.15	.51	1.05	.01	.99	18.21

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 4 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 16
Schmid and Leiman (1957) Solution for
Combined (n=1041) Sample for Ages 13 to 42 Months

Variable	Ī	II	III	A	В	С	D	h^2
1 Posit Affect	.024	.578			034	104	183	.471
2 Negat Affect	.065	.100	.686		054		059	.499
3 Soothability	.224	.075				.024	.000	.299
4 Hypersensiti	709	.062	.323		.358	045	019	.742
5 Energy	.159	.689	020	.009			193	.547
6 Adapt Change	.038	.017	758	075		019	.030	.583
7 Interest Tes	.148	.569		041		.020	146	.638
8 Initiative	.290	.612	416	031	.005	.174	136	.682
9 Exploration	.141	.752	072	.005			213	.642
10 Atten Tasks	.127	.277	738	069	.033	.060	036	.648
11 Persistence	.189	.390	662	059	.020	.094	069	.644
12 Enthusiasm	.039	.511	691	060	.033	027	125	.761
13 Fearfulness	690	164		005		039	.062	.645
14 Frus Inabili	593	.124	.358	.028	.277	067	049	.579
15 Orient Exami	030	.331	712	064	003	133	090	.648
16 Soc Engageme	.096	.615	233	011	056		 193	.487
17 Cooperation	.039	.240	823	078	.030	029	037	.746
18 Gross-motor	.563	.333	.044	.007	.034	.527		
19 Fine-motor	.596	.338	054	002	.019	.521	008	.743
20 Control Move	.640	.310	011	.004	036	.483		
21 Hypotonicity	768	.055	.025	012	.463	.030		
22 Hypertonicit	701	.063	.052	009	.454			
23 Tremulousnes	795	.065						
24 Slow Delayed	690	016	.010	014	.433	.056		
25 Frenetic Mov	783	.127	.144	.002				
26 Hyperactivit	664	.214	.335	.026	.324	066	072	.714
Trace	5.93	3.57	4.90	.04	1.46	.90	.24	17.04

Note. The column after the orthogonalized matrix presents the sum of the squared entries in a given row. The first 3 columns represent the second order factors. The next 4 columns represent the first order solution, based on variance orthogonal to the second order (Gorsuch, 1983, pp. 248-254).



Table 17 Mean Factor Score Differences Across Standardization and Clinical Samples As Regards the Varimax-Rotated Product $(F_{VxF}xF_{FxS} = F_{VxS})$ Matrices

Age 1 to 5 Months Factor Standardization Clinical **PCALCULATED** +.07 (1.10) +.58 (1.42) $1.\overline{19}$ I -.03 (0.95) -.26(0.57)99.16 <.0001 II Age 6 to 12 Months Factor Standardization Clinical PCALCULATED . 5304 +.02 (0.88) -.04 (1.17) <1.00 -.32 (0.28) +.55 (1.44) 108.89 I -.32 (0.28) -.02 (0.98) <.0001 ΙI +.03 (1.02) <1.00 .6101 Age 13 to 42 Months Clinical Factor Standardization PCALCULATED <.0001 203.11 +.32 (0.54) -.51 (1.31) -.02 (0.95) +.03 (1.08) -.02 (1.00) +.04 (1.00) I II <1.00 <1.00 .5009 .3630 III



